



# Probability and Statistics

## Tail Inequalities

谢润烁 Nanjing University, 2023 Fall



# Roadmap

- Tail Inequalities: Part I
  - Markov Inequality
  - Chebyshev Inequality
- Concentration of Measure
- Tail Inequalities: Part II
  - Chernoff Bounds
- Application
  - Randomized Quick Sort
  - The Median Trick
  - Load Balancing Problem
- More General Bounds

# Tail Inequalities: Part I

# In Analysis of Algorithms

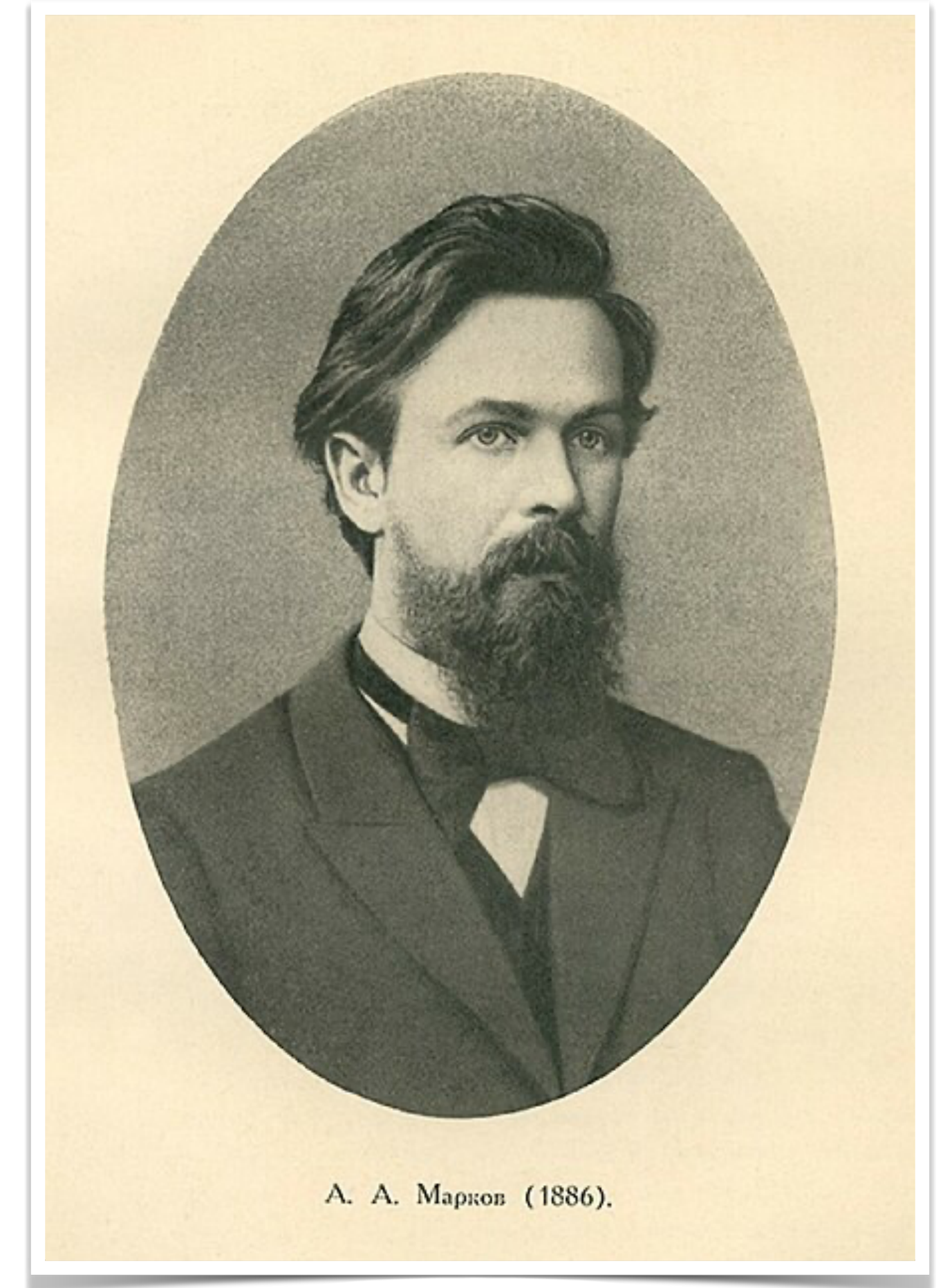
- Question:  $X$  is the running time of algorithm  $\mathcal{A}$ 
  - Is it possible that  $\Pr(X \geq \mathbb{E}X + t)$  very large?
- In analyzing the performance of a randomized algorithm, we often like to show that the behavior of the algorithm is good almost all the time
  - *i.e.* Establish high probability bounds on their run-time
  - *i.e.* Estimate the failure probability of algorithms

# Markov's Inequality

- For any random variable  $X \geq 0$  and  $a > 0$

$$\Pr(X \geq a) \leq \frac{\mathbb{E}[X]}{a}$$

- Gives the best tail bound possible when all we know is
  - the expectation of the random variable and
  - the variable is nonnegative



1856 ~ 1922

Андрей Андреевич Марков  
Andrey Andreyevich Markov

# Markov's Inequality

- For any random variable  $X \geq 0$  and  $a > 0$

$$\Pr(X \geq a) \leq \frac{\mathbb{E}[X]}{a}$$

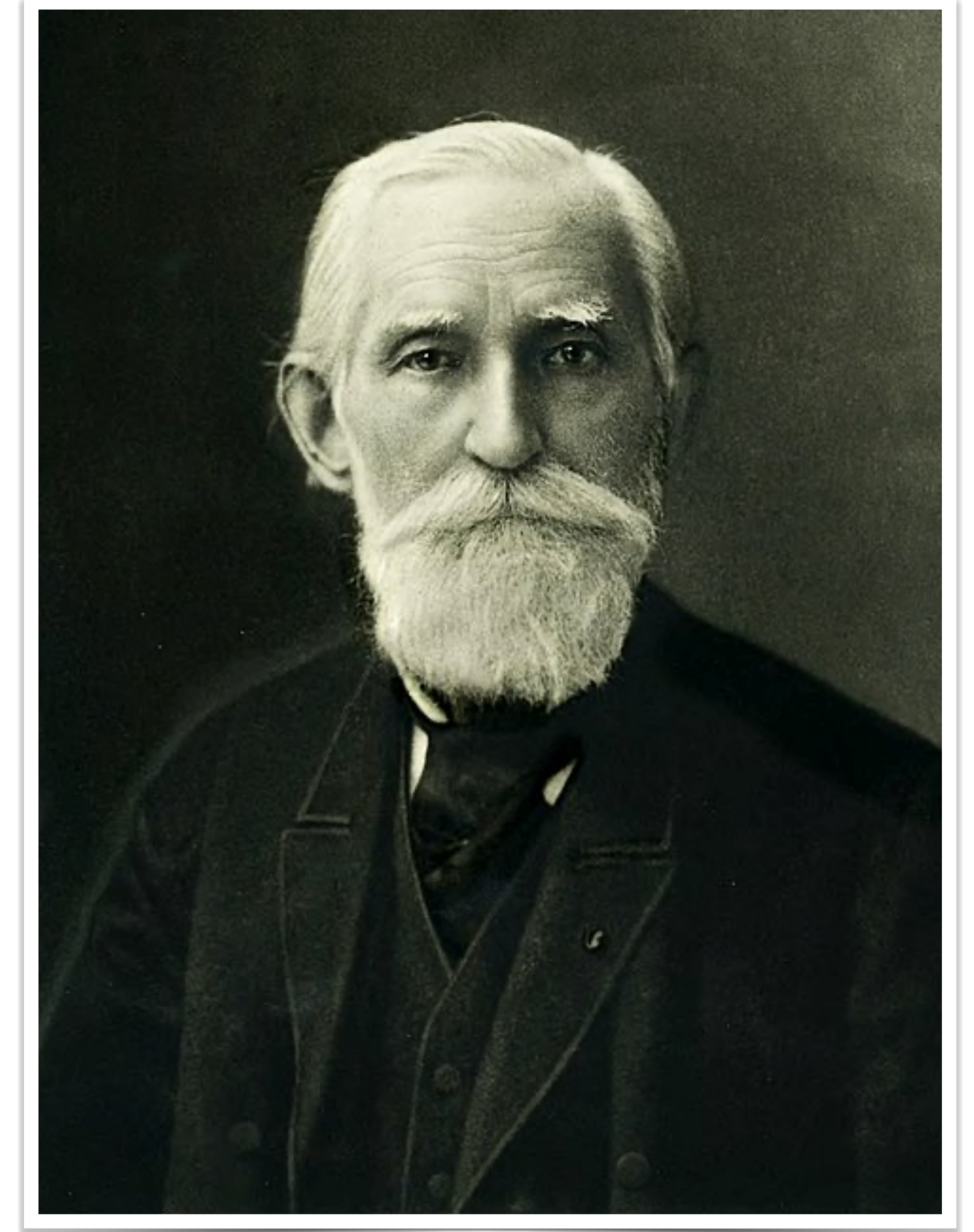
- Unfortunately, the Markov inequality is often too weak to yield useful results
- Can we leverage more information to gain a better bound?

# Chebyshev's Inequality

- For any  $a > 0$

$$\Pr(|X - \mathbb{E}[X]| \geq a) \leq \frac{\text{Var}[X]}{a^2}$$

- Proof: Apply Markov's Inequality



1821 ~ 1894

Пафну́тий Льво́вич Чебышёв  
Pafnuty Lvovich Chebyshev  
[pɐfˈnutʲɨj ˈlʲvovʲɪtɕ tɕɪbʲɪˈʂɔf]



# Chebyshev's Inequality

- For any  $a > 0$

$$\Pr(|X - \mathbb{E}[X]| \geq a) \leq \frac{\text{Var}[X]}{a^2}$$

- Is it better than Markov's Inequality?



# Chebyshev's Inequality

- For any  $a > 0$

$$\Pr(|X - \mathbb{E}[X]| \geq a) \leq \frac{\text{Var}[X]}{a^2}$$

- Is it better than Markov's Inequality?
- Consider flipping coins  $X \sim \text{Bin}(n, \frac{1}{2})$ 
  - $\Pr(X \geq \frac{3}{4}n)$

# Generalized Markov's Inequality

- For any  $f : \mathbb{R} \rightarrow \mathbb{R}_{\geq 0}$  and  $a > 0$

$$\Pr(f(X) \geq a) \leq \frac{\mathbb{E}[f(X)]}{a}$$

- Useful if  $f(X)$  can “extract” useful information about  $X$



# Generalized Markov's Inequality

- For any  $f : \mathbb{R} \rightarrow \mathbb{R}_{\geq 0}$  and  $a > 0$

$$\Pr(f(X) \geq a) \leq \frac{\mathbb{E}[f(X)]}{a}$$

- Example

- $k$ th moment method:

- $f(X) = \mathbb{E}[X^k]$   $k$ th moment

- $f(X) = \mathbb{E}[(X - \mathbb{E}X)^k]$   $k$ th central moment

- Chebyshev's inequality:  $f(X) = \mathbb{E}[(X - \mathbb{E}X)^2] = \text{Var}(X)$

- Chernoff-Hoeffding bounds:  $f(X) = \mathbb{E}[e^{\lambda X}]$

# Example: Weierstrass's Approximation Theorem

- Let  $f: [0,1] \rightarrow [0,1]$  be a continuous function. For any  $\epsilon > 0$ , there exists a polynomial such that

$$\sup_{x \in [0,1]} |p(x) - f(x)| \leq \epsilon$$

- For  $x \in [0,1]$ , let  $Y_x \sim \text{Bin}(n, x)$

$$p(x) = \mathbb{E}\left[f\left(\frac{Y_x}{n}\right)\right] \text{ is it a polynomial?}$$



# Example: Weierstrass's Approximation Theorem

- $|p(x) - f(x)| = \left| \mathbb{E} \left[ f\left(\frac{Y_x}{n}\right) - f(x) \right] \right| \leq \mathbb{E} \left[ \left| f\left(\frac{Y_x}{n}\right) - f(x) \right| \right]$
- Recall that
  - $f(x)$  is continuous in  $[a, b] \Rightarrow f(x)$  is uniformly continuous in  $[a, b]$
  - uniformly continuous  $\Rightarrow \exists \delta > 0$  s.t.  $|f(x) - f(y)| \leq \frac{\epsilon}{2}$  for all  $|x - y| \leq \delta$

# Example: Weierstrass's Approximation Theorem

$$|p(x) - f(x)| \leq \mathbb{E} \left[ \left| f\left(\frac{Y_x}{n}\right) - f(x) \right| \right]$$

- Denote  $A$  as event  $\left| \frac{Y_x}{n} - x \right| \leq \delta$

- Use conditional expectation

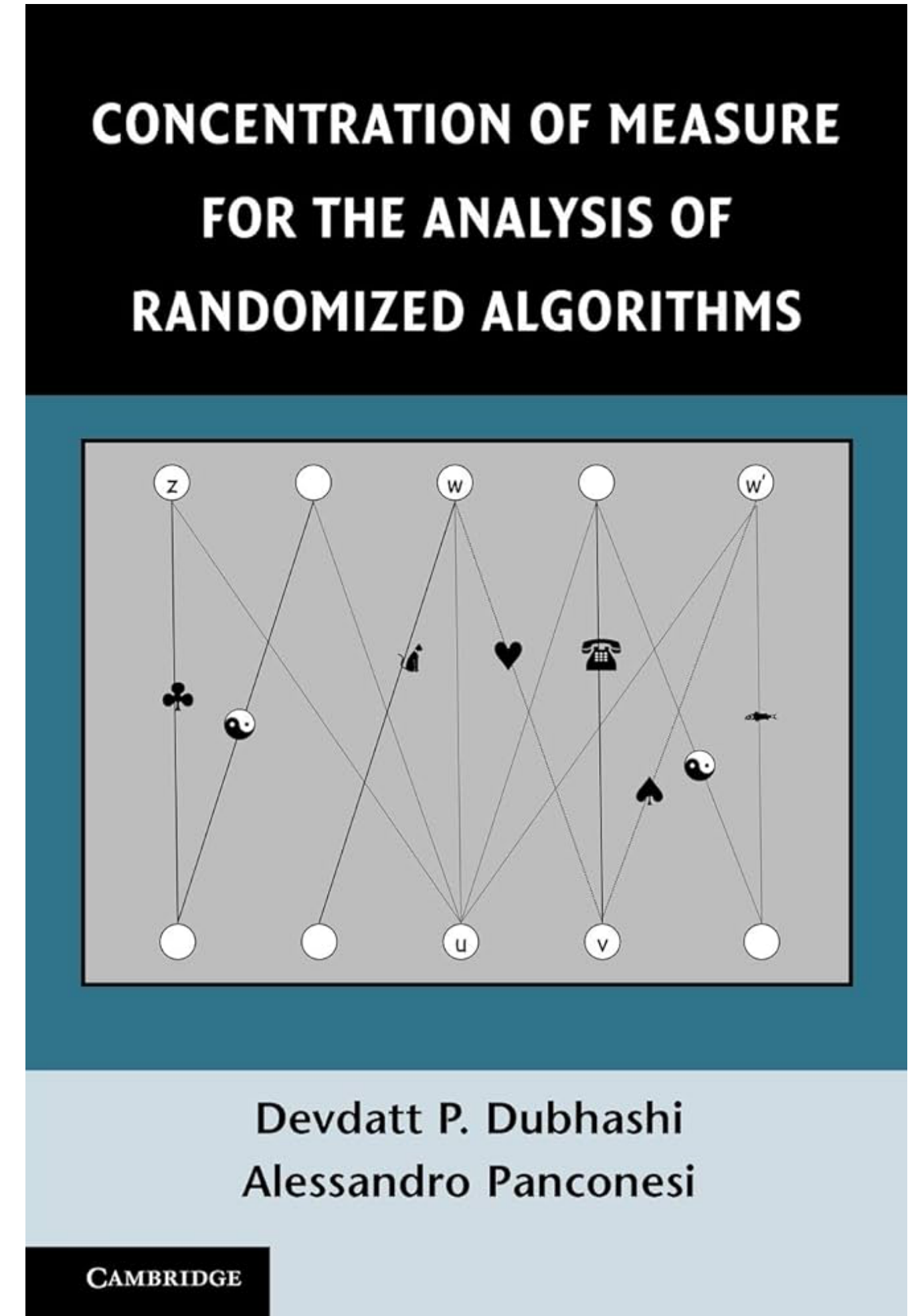
$$\begin{aligned} \mathbb{E} \left[ \left| f\left(\frac{Y_x}{n}\right) - f(x) \right| \right] &= \mathbb{E} \left[ \left| f\left(\frac{Y_x}{n}\right) - f(x) \right| \mid A \right] \Pr(A) + \mathbb{E} \left[ \left| f\left(\frac{Y_x}{n}\right) - f(x) \right| \mid A^c \right] \Pr(A^c) \\ &\leq \frac{\epsilon}{2} + \frac{1}{4n\delta^2} \leq \epsilon \quad \text{if } n \geq \frac{1}{2\epsilon\delta^2} \end{aligned}$$



# Concentration of Measure

# What is *Concentration of Measure*?

- The phenomenon that a function of a **large number** of random variables tends to concentrate its values in a **relatively narrow range**.
- Views from different scale
  - Micro: random
  - Macro: regular



# Law of Large Number

## The weak version, aka Khinchin(Хінчин)'s Law

- $X_1, X_2, \dots, X_n$  are independent and identically distributed random variables with mean  $\mu$  and standard deviation  $\sigma$
- For any constant  $\varepsilon > 0$  we have

$$\lim_{n \rightarrow \infty} \Pr \left( \left| \frac{\sum_{i=1}^n X_i}{n} - \mu \right| > \varepsilon \right) = 0$$

- You are able to prove this now



# Central Limit Theorem

- $X_1, X_2, \dots, X_n$  are independent and identically distributed random variables with mean  $\mu$  and standard deviation  $\sigma$
- For any real numbers  $a$  and  $b$  with  $a < b$ ,

$$\lim_{n \rightarrow \infty} \Pr \left( a \leq \frac{\sum_{i=1}^n X_i - n\mu}{\sigma\sqrt{n}} \leq b \right) = \Phi(b) - \Phi(a)$$

- Where  $\Phi(x)$  is the distribution function of standard normal distribution  $N(0,1)$

# Central Limit Theorem

$$p_1 = \frac{1}{6} \quad p_2 = \frac{1}{6} \quad p_3 = \frac{1}{6} \quad p_4 = \frac{1}{6} \quad p_5 = \frac{1}{6} \quad p_6 = \frac{1}{6}$$

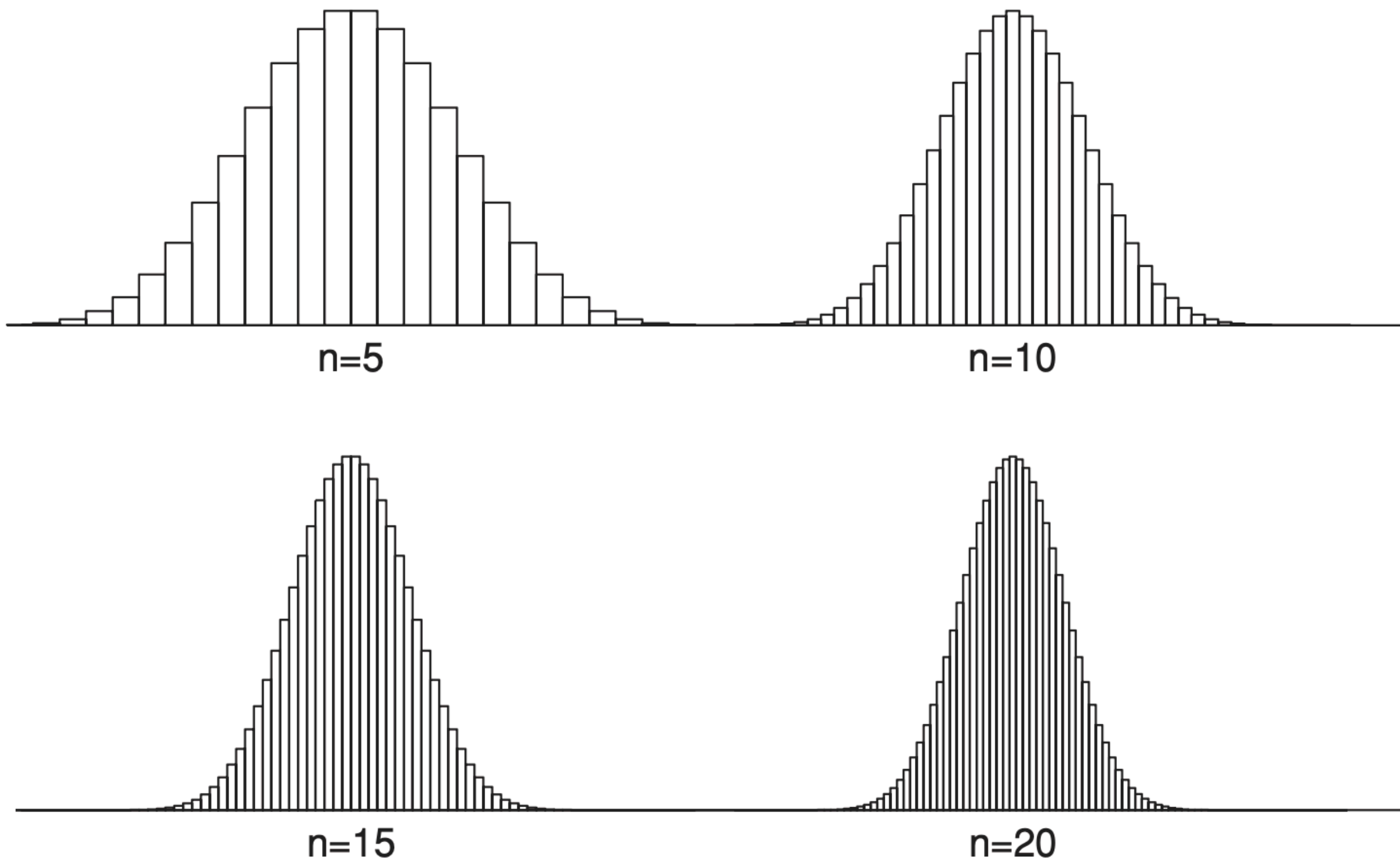


Fig. 5.6. Probability histogram for the unbiased die.

$$p_1 = 0.2 \quad p_2 = 0.1 \quad p_3 = 0.0 \quad p_4 = 0.0 \quad p_5 = 0.3 \quad p_6 = 0.4$$

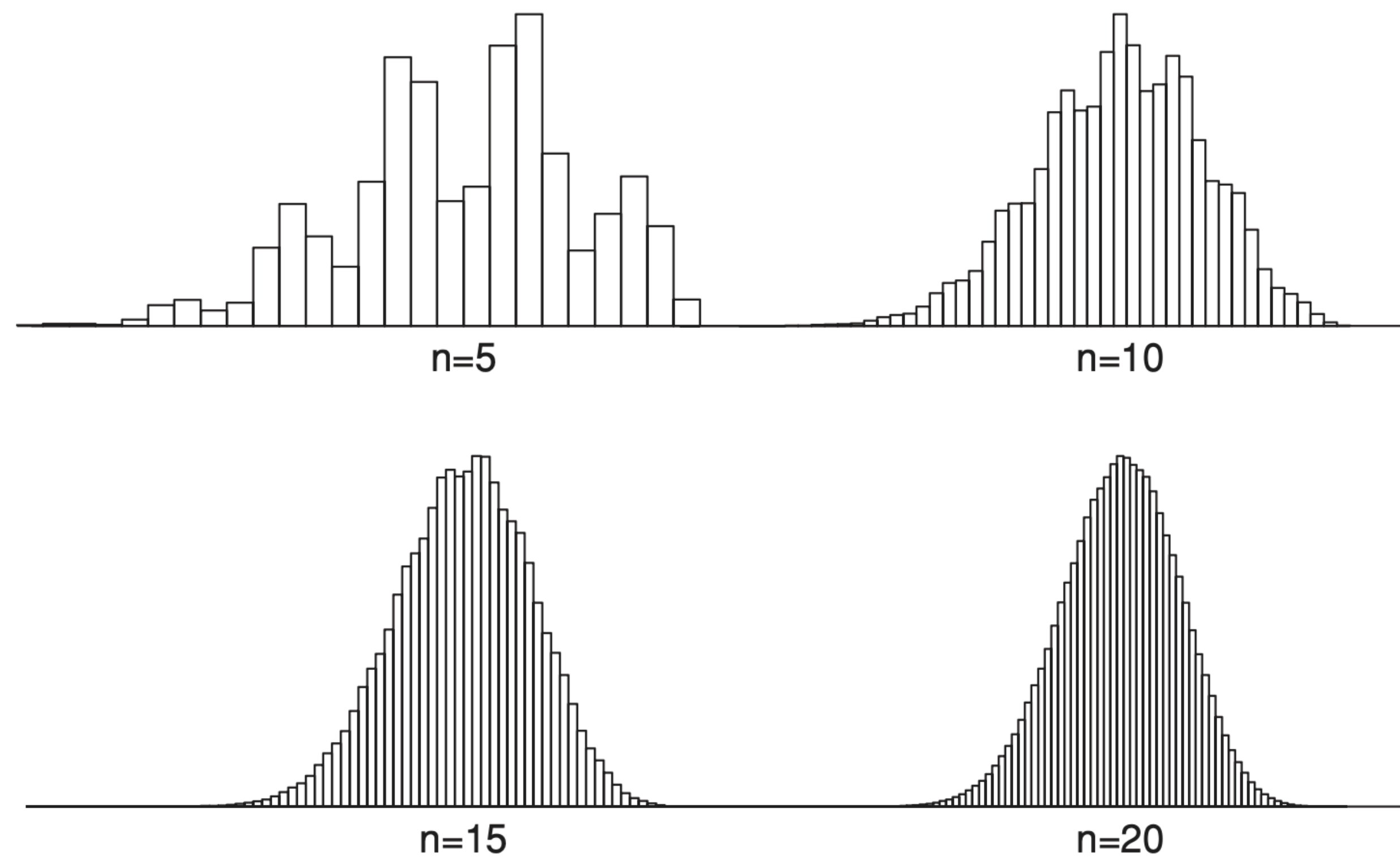


Fig. 5.7. Probability histogram for a biased die.

# In Analysis of Algorithms

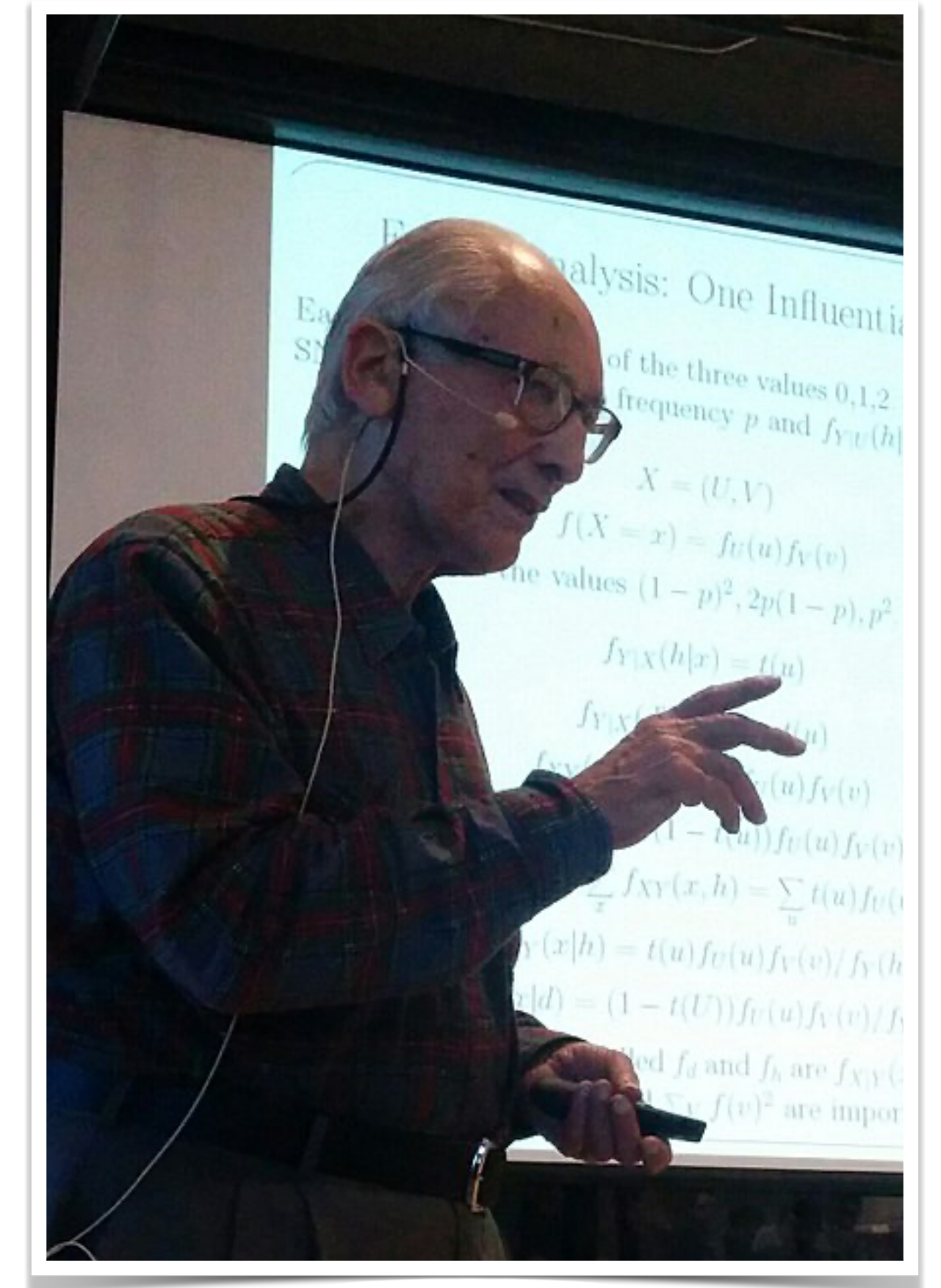
- These results (LLN, CLT) are **asymptotic** and **qualitative**
  - $n \rightarrow \infty$
- However, in the analysis of algorithms, we typically require **quantitative** estimates that are valid for finite (though large) values of  $n$

# Tail Inequalities: Part II



# Chernoff and Hoeffding Bounds

- Extremely powerful in analysis of algorithms
- Giving exponentially decreasing bounds on the tail distribution
- Derived by applying Markov's inequality to the moment generating function of a random variable



1923 ~  
Herman Chernoff

# Moment Generating Function (MGF)

- Moment Generating Function

$$M_X(\lambda) = \mathbb{E}[e^{\lambda X}]$$

- In many cases, the function is well-defined in the neighborhood of zero
- *Why Moment Generating?*

$$\mathbb{E}[X_n] = M_X^{(n)}(0)$$

# Chernoff Bounds

## Tight Forms

Let  $X = \sum_{i=1}^n X_i$  where  $X_1, X_2, \dots, X_n \in \{0,1\}$  are independent variables\*

Let  $\mu = \mathbb{E}[X]$

- for any  $\delta \geq 0$ ,

$$\Pr[X \geq (1 + \delta)\mu] \leq \left( \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right)^\mu \text{ (the Upper Tail)}$$

- for any  $0 \leq \delta \leq 1$ ,

$$\Pr[X \leq (1 - \delta)\mu] \leq \left( \frac{e^{-\delta}}{(1 - \delta)^{(1-\delta)}} \right)^\mu \text{ (the Lower Tail)}$$

# Proof Idea

## On Upper Tail

- $\Pr[X \geq (1 + \delta)\mu] = \Pr[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}] \leq \frac{\mathbb{E}[e^{\lambda X}]}{e^{\lambda(1+\delta)\mu}}$
- Find a  $\lambda$  to minimize  $\frac{\mathbb{E}[e^{\lambda X}]}{e^{\lambda(1+\delta)\mu}}$



# Chernoff Bounds

## Useful Forms

- For any  $0 < \delta < 1$ ,

$$\Pr[X \geq (1 + \delta)\mu] \leq \exp\left(-\frac{\mu\delta^2}{3}\right)$$

$$\Pr[X \leq (1 - \delta)\mu] \leq \exp\left(-\frac{\mu\delta^2}{2}\right)$$

- For  $t \geq 2e\mu$ ,

$$\Pr[X \geq t] \leq 2^{-t}$$

# Chernoff Bounds

- Compared to Markov's and Chebyshev's Inequalities
  - How is Chernoff Bounds' performance?
- Consider flipping coins  $X \sim \text{Bin}(n, \frac{1}{2})$  again
  - $\Pr(X \geq \frac{3}{4}n)$

**Application**

# The Median Trick

- Suppose we want to estimate the value of  $m$
- Let  $\mathcal{A}$  be an algorithm that outputs  $\hat{Z}$  satisfying

$$\Pr[(1 - \epsilon)m \leq \hat{Z} \leq (1 + \epsilon)m] \geq \frac{3}{4}$$

- How to improve our accuracy using  $\mathcal{A}$ ?
- Let  $X$  be the median of  $\hat{Z}_1, \hat{Z}_2, \dots, \hat{Z}_n$

$$\Pr[(1 - \epsilon)m \leq X \leq (1 + \epsilon)m] \geq ?$$



# Randomized Quicksort

- We denote  $X$  as the running time of randomized quicksort, *i.e.*, #comparisons
  - You've learned in your DS course that
  - $\mathbb{E}(X) = \Theta(n \log n)$

# Randomized Quicksort



## Randomized QuickSort

- Harmonic series

$$\triangleright H_n = \sum_{k=1}^n \frac{1}{k} \sim \ln n$$

- Hence,  $\mathbb{E}[X] < \sum_{i=1}^{n-1} \sum_{k=1}^n \frac{2}{k} < 2nH_n < 2n(1 + \ln n) = O(n \lg n)$

- Combined the fact that in the best case (balanced partition each time) randomized quick sort is  $\Theta(n \lg n)$ , the expected running time is  $\Theta(n \lg n)$ .
- In fact, runtime of `RndQuickSort` is  $O(n \log n)$  with high probability!

# Randomized Quicksort

- Now we can prove that the running time is  $O(n \lg n)$  with high probability

$$\text{i.e. } \lim_{n \rightarrow \infty} \Pr[X > O(n \lg n)] = 0$$

- Can we use the way we analyze the expected running time?

# Load Balancing / Occupancy

## Balls into Bins Model

- We throw  $m$  balls into  $n$  bins uniformly and independently
- $Y_i$ : number of balls, which is called the load, in the  $i$ -th bin

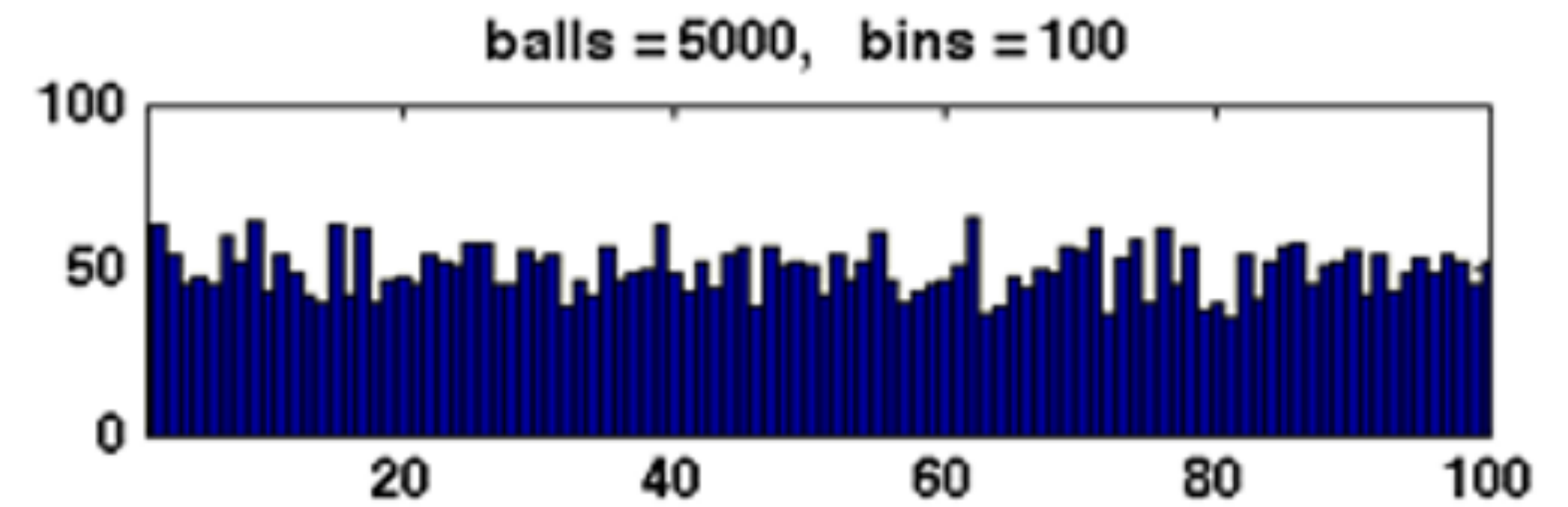
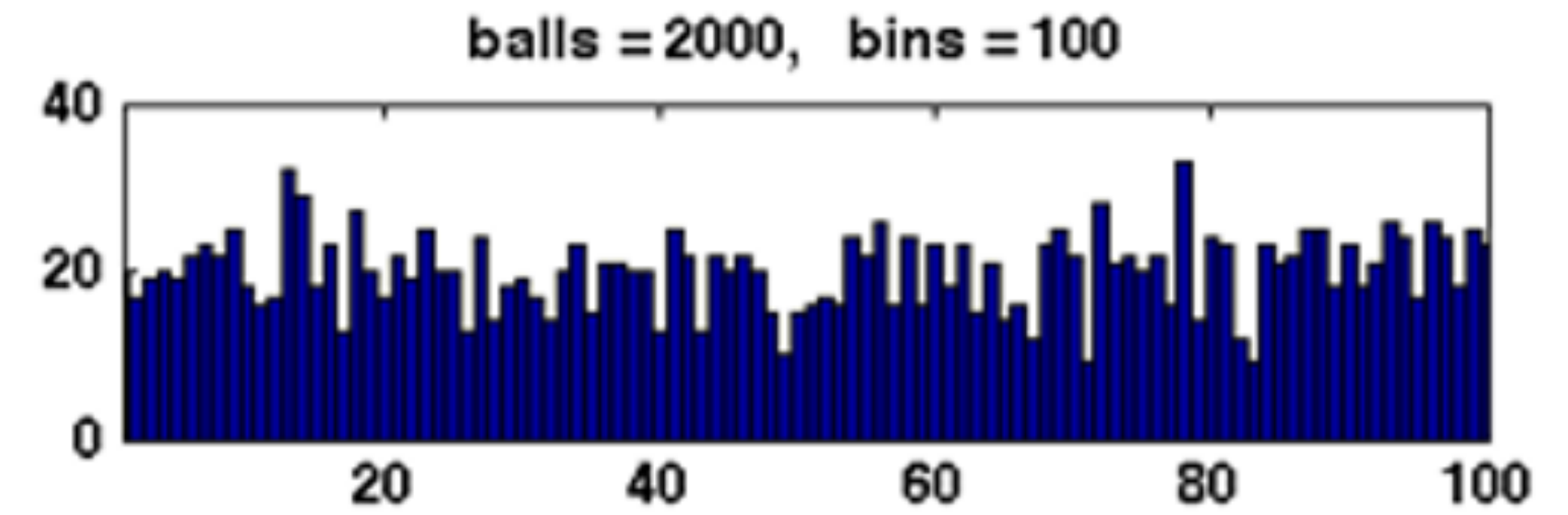
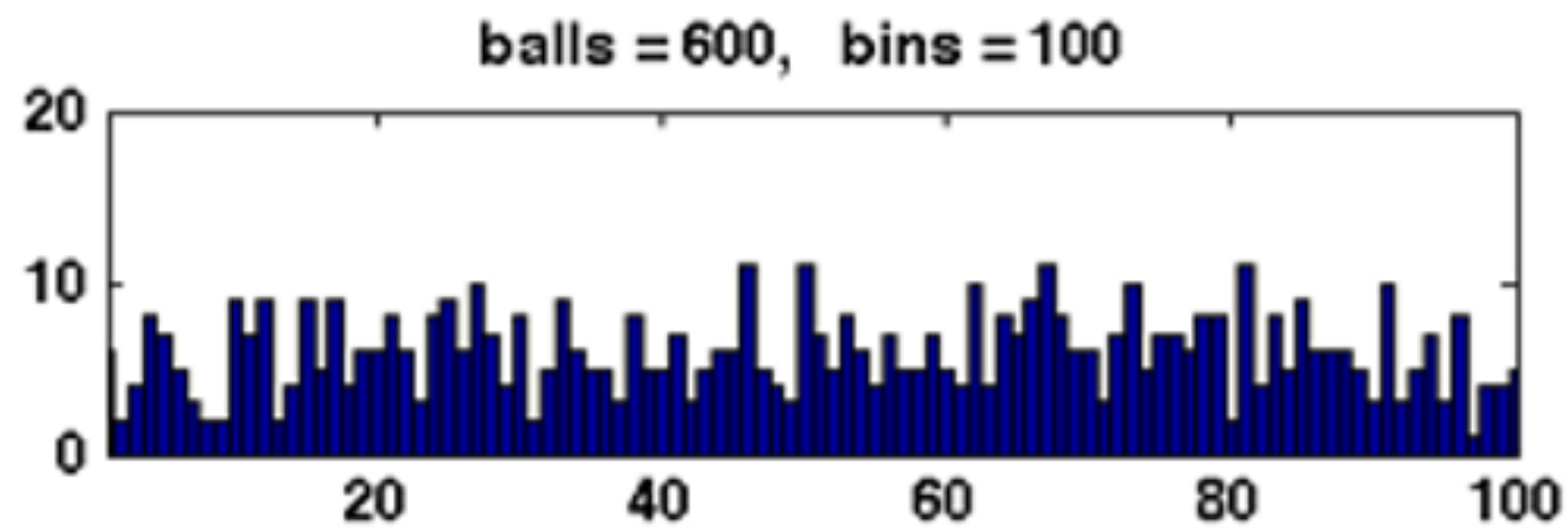
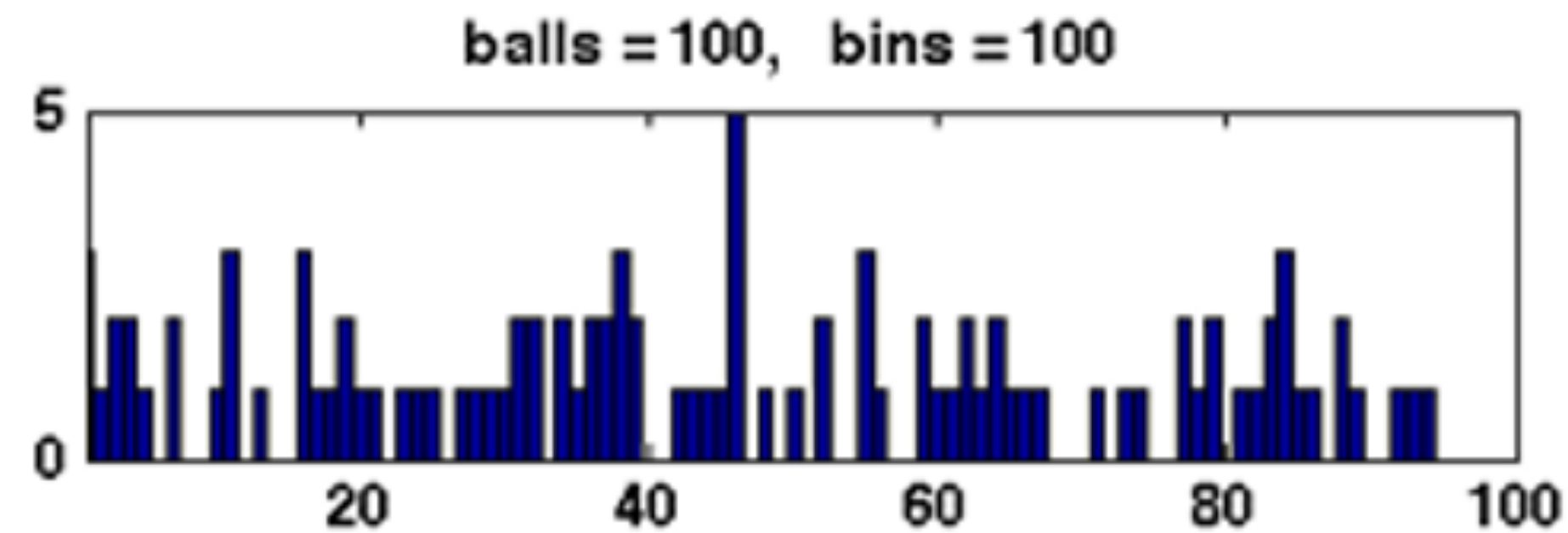
$$\mathbb{E}(Y_i) = \frac{m}{n}$$

- What is the maximum load of all bins?



# Load Balancing / Occupancy

## Balls into Bins Model



# Load Balancing / Occupancy

## Balls into Bins Model

- When  $m = n$ , the maximum load is

$$O\left(\frac{e \ln n}{\ln \ln n}\right) \text{ w.h.p.}$$

- When  $m > n \ln n$ , the maximum load is

$$O\left(\frac{m}{n}\right) \text{ w.h.p.}$$

# More General Bounds

# Chernoff-Hoeffding Bounds

- Let  $X_1, \dots, X_n$  be independent random variables with  $\Pr(a_i \leq X_i \leq b_i) = 1$  for constants  $a_i$  and  $b_i$ . Then

$$\Pr\left(|X - \mu| \geq \varepsilon\right) \leq 2e^{\frac{-2\varepsilon^2}{\sum_{i=1}^n (b_i - a_i)^2}}$$

- Where  $X = \sum_{i=1}^n X_i$ ,  $\mu = \mathbb{E}(X) = \sum_{i=1}^n \mathbb{E}[X_i]$

# The Method of Bounded Differences

- For independent  $X_1, \dots, X_n$ , if  $n$ -variate function  $f$  satisfies the **Lipschitz condition**: for every  $1 \leq i \leq n$  and all  $x_1, \dots, x_n$  and  $y_i$

$$\left| f(x_1, \dots, x_n) - f(x_1, \dots, x_{i-1}, y_i, x_{i+1}, \dots, x_n) \right| \leq c_i$$

- Then for any  $\epsilon > 0$ :

$$\Pr \left[ \left| f(X_1, \dots, X_n) - \mathbb{E}(f(X_1, \dots, X_n)) \right| \geq \epsilon \right] \leq 2e^{\frac{-2\epsilon^2}{\sum_{i=1}^n c_i}}$$